

Package ‘glmmAK’

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Author Arnošt Komárek <arnost.komarek@mff.cuni.cz>

Maintainer Arnošt Komárek <arnost.komarek@mff.cuni.cz>

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Description Later

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URL <http://www.karlin.mff.cuni.cz/~komarek>

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| AKmiscel | <i>Miscellaneous smaller functions</i> |
|----------|----------------------------------------|

Description

Functions `mfLand` and `mfPort` try to split a graphical device in a “nice” way to produce several plots on 1 page in landscape or portrait format.

Usage

`mfLand(np)`

`mfPort(np)`

Arguments

`np` number of plots that are to be produced on 1 page

Value

A 2-component vector giving the number of rows and columns into which the graphical device could be splitted.

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

See Also

[par.](#)

Examples

`mfLand(6)`

`mfPort(6)`

BPvalue*Univariate Bayesian P-value (pseudo-contour probability)*

Description

For a sample (from the posterior distribution) of θ this function computes

$$P = 2 \min\{P(\theta < 0), P(\theta > 0)\},$$

which can be viewed as a counterpart of a classical two-sided P-value.

Note that this is the same as a univariate pseudo-contour probability as described in Besag et al. (1995, p. 30) and in Held (2004).

Usage

```
BPvalue(sample)
```

Arguments

| | |
|--------|-------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| sample | vector, matrix or data frame with sampled values. If it is a matrix or data frame then it is assumed that each column corresponds to a separate parameter and the P-value is computed separately for each column. |
|--------|-------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|

Value

Vector of computed P-values.

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

References

Besag, J., Green, P., Higdon, D. and Mengersen, K. (1995). Bayesian computation and stochastic systems (with Discussion). *Statistical Science*, **10**, 3 - 66.

Held, L. (2004). Simultaneous posterior probability statements from Monte Carlo output. *Journal of Computational and Graphical Statistics*, **13**, 20 - 35.

Examples

```
m <- 1000
sample <- rnorm(m, mean=1)
BPvalue(sample)
## compare with
2*pnorm(0, mean=1, lower.tail=TRUE)

sample <- data.frame(x1=rnorm(m), x2=rnorm(m, mean=-1), x3=rnorm(m, mean=2))
BPvalue(sample)
```

```
## compare with
2*pnorm(0, mean=0)
2*pnorm(0, mean=-1, lower.tail=FALSE)
2*pnorm(0, mean=2, lower.tail=TRUE)
```

 copula

Copulas

Description

Functions to compute the cumulative distribution functions and densities for several bivariate copulas.

These functions do not have anything to do with the GLMM's in this package. They are here simply because of an interest of the author to play with copulas a little bit.

Usage

```
Cplackett(u, v, theta=1)
Cgauss(u, v, theta=0)
Cclayton(u, v, theta=0)
cplackett(u, v, theta=1)
cgauss(u, v, theta=0)
cclayton(u, v, theta=0)
```

Arguments

| | |
|-------|----------------------------------------------------------------------------|
| u | Unif(0, 1) quantiles for the first margin. It can be a vector or a matrix |
| v | Unif(0, 1) quantiles for the second margin. It can be a vector or a matrix |
| theta | value of the association parameter |

Value

A vector or a matrix with the values of the corresponding cdf or density.

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

References

Nelsen, R. B. (2006). *An Introduction to Copulas*, Second Edition New York: Springer.

Examples

```

### Margin1 = normal mixture
### Margin2 = normal
intcpt <- c(1, 5)
sds <- c(1, 3)
x <- seq(intcpt[1]-2.5*sds[1], intcpt[1]+1+2.5*sds[1], length=40)
y <- seq(intcpt[2]-2.5*sds[2], intcpt[2]+2.5*sds[2], length=41)
Fx <- 0.6*pnorm(x, mean=intcpt[1], sd=sds[1]) +
      0.4*pnorm(x, mean=intcpt[1]+2, sd=0.5*sds[1])
Fy <- pnorm(y, mean=intcpt[2], sd=sds[2])
fx <- 0.6*dnorm(x, mean=intcpt[1], sd=sds[1]) +
      0.4*dnorm(x, mean=intcpt[1]+2, sd=0.5*sds[1])
fy <- dnorm(y, mean=intcpt[2], sd=sds[2])
u <- matrix(rep(Fx, length(y)), ncol=length(y))
v <- matrix(rep(Fy, length(x)), nrow=length(x), byrow=TRUE)
du <- matrix(rep(fx, length(y)), ncol=length(y))
dv <- matrix(rep(fy, length(x)), nrow=length(x), byrow=TRUE)

### Copula distribution functions
theta <- c(3, 0.3, 1)
CC <- list()
CC$plackett <- Cplackett(u, v, theta=theta[1])
CC$gauss <- Cgauss(u, v, theta=theta[2])
CC$cclayton <- Cclayton(u, v, theta=theta[3])

### Copula densities
cc <- list()
cc$plackett <- cplackett(u, v, theta=theta[1]) * du * dv
cc$gauss <- cgauss(u, v, theta=theta[2]) * du * dv
cc$cclayton <- ccclayton(u, v, theta=theta[3]) * du * dv

### Figures
lcol <- "red"
pcol <- "seagreen2"
mains <- paste(c("Plackett ", "Gauss ", "Clayton "), "copula, theta=", theta, sep="")
zlab <- "F(x,y)"
zlab2 <- "f(x,y)"
tangle <- -25
ptangle <- 40

oldpar <- par(bty="n", mfcol=c(2, 3))
contour(x, y, cc$plackett, main=mains[1], col=lcol)
persp(x, y, cc$plackett, zlab=zlab2, main=mains[1], col=pcol, theta=tangle, phi=ptangle)
contour(x, y, cc$gauss, main=mains[2], col=lcol)
persp(x, y, cc$gauss, zlab=zlab2, main=mains[2], col=pcol, theta=tangle, phi=ptangle)
contour(x, y, cc$cclayton, main=mains[3], col=lcol)
persp(x, y, cc$cclayton, zlab=zlab2, main=mains[3], col=pcol, theta=tangle, phi=ptangle)

par(bty="n", mfcol=c(2, 3))
contour(x, y, CC$plackett, main=mains[1], col=lcol)
persp(x, y, CC$plackett, zlab=zlab, main=mains[1], col=pcol, theta=tangle, phi=ptangle)

```

```

contour(x, y, CC$gauss, main=mains[2], col=lcol)
persp(x, y, CC$gauss, zlab=zlab, main=mains[2], col=pcol, theta=tangle, phi=ptangle)
contour(x, y, CC$clayton, main=mains[3], col=lcol)
persp(x, y, CC$clayton, zlab=zlab, main=mains[3], col=pcol, theta=tangle, phi=ptangle)

par(bty="n")
layout(matrix(c(0,1,1,0, 2,2,3,3), nrow=2, byrow=TRUE))
contour(x, y, cc$plackett, main=mains[1], col=lcol)
contour(x, y, cc$gauss, main=mains[2], col=lcol)
contour(x, y, cc$clayton, main=mains[3], col=lcol)

par(bty="n")
layout(matrix(c(0,1,1,0, 2,2,3,3), nrow=2, byrow=TRUE))
persp(x, y, cc$plackett, zlab=zlab2, main=mains[1], col=pcol, theta=tangle, phi=ptangle)
persp(x, y, cc$gauss, zlab=zlab2, main=mains[2], col=pcol, theta=tangle, phi=ptangle)
persp(x, y, cc$clayton, zlab=zlab2, main=mains[3], col=pcol, theta=tangle, phi=ptangle)

par(bty="n")
layout(matrix(c(0,1,1,0, 2,2,3,3), nrow=2, byrow=TRUE))
contour(x, y, CC$plackett, main=mains[1], col=lcol)
contour(x, y, CC$gauss, main=mains[2], col=lcol)
contour(x, y, CC$clayton, main=mains[3], col=lcol)

par(bty="n")
layout(matrix(c(0,1,1,0, 2,2,3,3), nrow=2, byrow=TRUE))
persp(x, y, CC$plackett, zlab=zlab, main=mains[1], col=pcol, theta=tangle, phi=ptangle)
persp(x, y, CC$gauss, zlab=zlab, main=mains[2], col=pcol, theta=tangle, phi=ptangle)
persp(x, y, CC$clayton, zlab=zlab, main=mains[3], col=pcol, theta=tangle, phi=ptangle)

par(oldpar)

```

cumlogit

Cumulative logit model for ordinal responses

Description

Fits the cumulative logit model using the maximum-likelihood. The log-likelihood is maximized using the Newton-Raphson algorithm. The function returns the inverse of both observed and expected information matrix. Summary of the model produced by the summary function uses by default the inverse of the observed information matrix for the inference. This can be changed by the user such that the expected information is used instead.

Usage

```

cumlogit(y, v, x, C=1, logit.order=c("decreasing", "increasing"),
         epsilon=1e-08, maxit=25, trace=FALSE)

## S3 method for class 'cumlogit'
print(x, vcov=c("observed", "expected"), ...)

```

```
## S3 method for class 'cumlogit'
summary(object, vcov=c("observed", "expected"), ...)
```

Arguments

y response vector taking values $0, 1, \dots, C$.

v matrix or data.frame with covarites whose effect does not necessarily satisfy proportional odds assumption. Intercept is included by default in the model and should be included neither in **x**, nor in **v**.

x vector, matrix or data.frame with covarites whose effect is assumed to satisfy proportional odds assumption.

C number of response categories minus 1.

logit.order either "decreasing" or "increasing" indicating in which direction the logits are formed.

For **logit.order**="decreasing":

$$\begin{aligned} \log \left\{ \frac{P(Y \geq 1)}{P(Y = 0)} \right\} &= \beta'x + \gamma'_1 v \\ \log \left\{ \frac{P(Y \geq 2)}{P(Y \leq 1)} \right\} &= \beta'x + \gamma'_2 v \\ &\vdots \\ \log \left\{ \frac{P(Y = C)}{P(Y \leq C-1)} \right\} &= \beta'x + \gamma'_C v \end{aligned}$$

For **logit.order**="increasing":

$$\begin{aligned} \log \left\{ \frac{P(Y = 0)}{P(Y \geq 1)} \right\} &= \beta'x + \gamma'_1 v \\ \log \left\{ \frac{P(Y \leq 1)}{P(Y \geq 2)} \right\} &= \beta'x + \gamma'_2 v \\ &\vdots \\ \log \left\{ \frac{P(Y \leq C-1)}{P(Y = C)} \right\} &= \beta'x + \gamma'_C v \end{aligned}$$

vcov character indicating which type of the information matrix should be used for the inference

epsilon positive convergence tolerance ε . The iterations converge when

$$\left| \frac{\ell_{new} - \ell_{old}}{\ell_{new}} \right| \leq \varepsilon,$$

where ℓ denotes the value of the log-likelihood.

maxit integer giving the maximal number of iterations.

trace logical indicating if output should be produced for each iteration.

object an object of class "cumlogit".

... other arguments passed to print or summary.

Value

An object of class "cumlogit". This has components

| | |
|-------------------|-------------------------------------------------------------------------|
| coefficients | the coefficients of the linear predictor. |
| loglik | the value of the log-likelihood. |
| score | the score vector. |
| vcov | the inverse of the observed information matrix. |
| expect.vcov | the inverse of the expected information matrix. |
| logit.order | character indicating the way in which the logits are formed. |
| linear.predictors | the values of the linear predictor for each observation and each logit. |
| fitted.values | the values of fitted category probabilities for each observation. |
| converged | logical indicating whether the optimization routine converged. |
| iter | number of iterations performed |
| C | see the function argument |
| y | see the function argument |
| v | see the function argument |
| x | see the function argument |

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

References

Agresti, A. (2002). *Categorical Data Analysis. Second edition.* Hoboken: John Wiley & Sons. Section 7.2.

See Also

[glm](#), [polr](#).

Examples

```
## Simulate the data for C = 3
## =====
set.seed(775988621)
N3 <- 1000

## covariates:
x3 <- data.frame(x1=rnorm(N3), x2=runif(N3, 0, 1))
v3 <- data.frame(v1=rnorm(N3), v2=runif(N3, 0, 1))

## regression coefficients
## (theta = c(0.1, -0.2, 1, 0.1, -0.2, 0, 0.05, -0.25, -0.5, 0.05, -0.25)):
alpha3 <- c(1, 0, -0.5)
```

```

beta3 <- c(0.1, -0.2)
gamma3 <- list(c(0.1, -0.2), c(0.05, -0.25), c(0.05, -0.25))

## linear predictors and inverse logits:
eta3 <- data.frame(eta1=alpha3[1] + as.matrix(x3) %*% beta3 +
  as.matrix(v3) %*% gamma3[[1]],
  eta2=alpha3[2] + as.matrix(x3) %*% beta3 +
  as.matrix(v3) %*% gamma3[[2]],
  eta3=alpha3[3] + as.matrix(x3) %*% beta3 +
  as.matrix(v3) %*% gamma3[[3]])
ilogit3 <- data.frame(ilogit1 = exp(eta3[,1])/(1 + exp(eta3[,1])),
  ilogit2=exp(eta3[,2])/(1 + exp(eta3[,2])),
  ilogit3=exp(eta3[,3])/(1 + exp(eta3[,3])))

## category probabilities:
pis3 <- data.frame(pi0=1-ilogit3[,1],
  pi1=ilogit3[,1]-ilogit3[,2],
  pi2=ilogit3[,2]-ilogit3[,3],
  pi3=ilogit3[,3])

## response:
rtemp <- function(prob, C){
  return(sample(0:C, size=1, prob=prob))
}
y3 <- apply(pis3, 1, rtemp, C=3)

## Fit the model
## =====
fit <- cumlogit(y=y3, x=x3, v=v3, C=3)
print(fit)

fit2 <- cumlogit(y=y3, x=x3, v=v3, C=3, logit.order="increasing")
print(fit2)

```

cumlogitRE

Logit and cumulative logit model with random effects

Description

This function implements MCMC sampling for the logit model with binary response and the cumulative logit model for multinomial ordinal response. Details are given in Komárek and Lesaffre (2008). On as many places as possible, the same notation as in this paper is used also in this manual page.

In general, the following (cumulative) logit model for response Y is assumed:

$$\begin{aligned}
 \log \left\{ \frac{P(Y \geq 1)}{P(Y = 0)} \right\} &= \eta^{(1)} \\
 \log \left\{ \frac{P(Y \geq 2)}{P(Y \leq 1)} \right\} &= \eta^{(2)} \\
 &\vdots \\
 \log \left\{ \frac{P(Y = C)}{P(Y \leq C-1)} \right\} &= \eta^{(C)},
 \end{aligned}$$

where the form of the linear predictors η_1, \dots, η_C depends on whether a hierarchical centering is used or not. In the following, β denotes fixed effects and b random effects.

No hierarchical centering (DEFAULT)

The linear predictor for the c th logit has the following form

$$\eta^{(c)} = \beta^{(c)'}(v', v'_b) + \beta^{*'}(x', x'_b) + b'(v'_b, x'_b) \quad (c = 1, \dots, C),$$

where $\beta = (\beta^{(1)'}, \dots, \beta^{(C)'}, \beta^{*'})'$ is the vector of the fixed-effects and b is a vector of random effects with zero location.

Hierarchical centering

The linear predictor for the c th logit has the following form

$$\eta^{(c)} = \beta^{(c)'}v + \beta^{*'}x + b^{(c)'}v_b + b^{*'}x_b \quad (c = 1, \dots, C),$$

where $\beta = (\beta^{(1)'}, \dots, \beta^{(C)'}, \beta^{*'})'$ is the vector of the fixed-effects and $b = (b^{(1)'}, \dots, b^{(C)'}, b^{*'})'$ is a vector of random effects with location $\alpha = (\alpha^{(1)'}, \dots, \alpha^{(C)'}, \alpha^{*'})'$.

Normal random effects (drandom="normal")

A vector of random effects is assumed to follow a (multivariate) normal distribution.

That is, if there is **no hierarchical centering** we assume for b :

$$b \sim N(\mathbf{0}, \mathbf{D}_b),$$

where \mathbf{D}_b is their variance-covariance matrix.

If the random effects are **hierarchically centered** then we assume for $b = (b^{(1)'}, \dots, b^{(C)'}, b^{*'})'$:

$$b \sim N(\alpha, \mathbf{D}_b),$$

where $\alpha = (\alpha^{(1)'}, \dots, \alpha^{(C)'}, \alpha^{*'})'$ is a vector of random effect locations (means) and \mathbf{D}_b is their variance-covariance matrix.

Further, in a Bayesian model, it is assumed that \mathbf{D}_b^{-1} has a Wishart $W(\nu_b, S_b)$ prior with ν_b degrees of freedom and scale matrix S_b . That is, a priori

$$\mathbf{D}_b^{-1} \sim W(\nu_b, S_b),$$

$$E(\mathbf{D}_b^{-1}) = \nu_b S_b.$$

Note that ν_b must be higher than the number of random effects minus 1.

Alternatively, when there is only a univariate random effect with the variance d_b^2 , it is possible to specify a uniform prior for the standard deviation of the random effect. That is, a priori

$$d_b \sim \text{Unif}(0, S).$$

G-spline distributed random effects (drandom="gspline")

See [gspline1](#) and [gspline2](#) for description of the G-spline (penalized Gaussian mixture) distribution. Further, see Komárek and Lesaffre (2008) for description of the prior distribution on the G-spline. Brief description follows here as well.

Univariate G-spline

If there is only a univariate random effect in the model and its distribution is specified as a G-spline that it is assumed that

$$b \sim \alpha + \sum_{j=-K}^K w_j N(\tau\mu_j, (\tau\sigma)^2),$$

where α is a location parameter, τ is a scale parameter and $w = (w_{-K}, \dots, w_K)'$ is a vector of G-spline weights. For **hierarchically centered** random effects, the location parameter α is fixed to zero.

Further, $M = (\mu_{-K}, \dots, \mu_K)'$ is a vector of fixed equidistant knots (component means), where

$$\mu_j = j\delta \quad (j = -K, \dots, K)$$

and σ is a fixed basis standard deviation.

The constraints $0 < w_j < 1$ ($j = -K, \dots, K$) and $\sum_{j=-K}^K w_j = 1$ are sufficient for G-spline to be a density. To avoid constraint estimation we will estimate transformed weights $a = (a_{-K}, \dots, a_K)'$ instead which relates to the original weights by

$$w_j = \frac{\exp(a_j)}{\sum_{k=-K}^K \exp(a_k)} \quad (j = -K, \dots, K),$$

$$a_j = \log \frac{w_j}{w_0} \quad (j = -K, \dots, K),$$

In the estimation procedure a penalty on the a coefficients in the form of a Gaussian Markov random field prior is imposed.

Bivariate G-spline

If there is a bivariate random effect $b = (b_1, b_2)'$ in the model and its distribution is specified as a G-spline that it is assumed that

$$b \sim (\alpha_1, \alpha_2)' + \sum_{j_1=-K_1}^{K_1} \sum_{j_2=-K_2}^{K_2} w_{j_1, j_2} \mathbf{N}\left((\tau_1 \mu_{1, j_1}, \tau_2 \mu_{2, j_2})', \text{diag}((\tau_1 \sigma_1)^2, (\tau_2 \sigma_2)^2)\right),$$

where α_1, α_2 are location parameters, τ_1, τ_2 are scale parameters and $W = (w_{-K_1, -K_2}, \dots, w_{K_1, K_2})'$ is a matrix of G-spline weights. For **hierarchically centered** random effects, the location parameters α_1, α_2 are fixed to zero.

Further, $M_1 = (\mu_{1, -K_1}, \dots, \mu_{1, K_1})'$ is a vector of fixed equidistant knots (component means) in the first margin, where

$$\mu_{1, j_1} = j_1 \delta_1 \quad (j_1 = -K_1, \dots, K_1)$$

and σ_1 is a fixed basis standard deviation in the first margin. Similarly for the second margin.

Similar reparametrization of the G-spline weights as in the univariate case is used to avoid constrained estimation.

Usage

```
cumlogitRE(y, v, x, vb, xb, cluster,
  intcpt.random=FALSE,
  hierar.center=FALSE,
  drandom=c("normal", "gspline"),
  C=1,
  logit.order=c("decreasing", "increasing"),
  prior.fixed,
  prior.random,
  prior.gspline,
```

```

init.fixed,
init.random,
init.gspline,
nsimul = list(niter=10, nthin=1, nburn=0, nwrite=10),
store = list(prob=FALSE, b=FALSE, alloc=FALSE, acoef=FALSE),
dir=getwd(),
precision=8)

```

Arguments

| | |
|---------------|---------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| y | response vector taking values $0, 1, \dots, C$. |
| v | vector, matrix or data.frame with covarites for fixed effects whose effect does not necessarily satisfy proportional odds assumption. If the argument <code>intcpt.random</code> is set to FALSE then the fixed intercept is included by default in the model. The intercept column should be included neither in x, nor in v. |
| x | vector, matrix or data.frame with covarites for fixed effects whose effect is assumed to satisfy proportional odds assumption. |
| vb | vector, matrix or data.frame with covarites for random effects whose effect does not necessarily satisfy proportional odds assumption. If you want to include random intercept , do it by setting the argument <code>intcpt.random</code> to TRUE. The intercept column should be included neither in xb, nor in vb. |
| xb | vector, matrix or data.frame with covarites for random effects whose effect is assumed to satisfy proportional odds assumption. |
| cluster | vector which determines clusters. Needed only when there are any random effects in the model. |
| intcpt.random | logical indicating whether a random intercept should be included in the model. |
| hierar.center | logical indicating whether a hierarchical centering of random effects should be used or not. |
| drandom | character indicating assumed distribution of random effects (if there are any). |
| C | number of response categories minus 1. |
| logit.order | either "decreasing" or "increasing" indicating in which direction the logits are formed. See the same argument in cumlogit for more details. Currently, only "decreasing" is implemented for <code>cumlogitRE</code> . |
| prior.fixed | list specifying the prior distribution for the fixed effects (regression coefficients). mean vector giving the prior mean for each regression coefficient. It can be a single number only in which case it is recycled. var vector giving the prior variance for each regression coefficient. It can be a single number only in which case it is recycled. In the case that prior mean and/or variances are different for different regression coefficients then they should be given in the following order: intercept for the first logit, $\beta(v)$ for the first logit, \dots , intercept for the last logit, $\beta(v)$ for the last logit, $\beta(x)$, |

`prior.random` list specifying the prior distribution for the parameters of the distribution of random effects. Composition of this list depends on the chosen distribution of random effects (**normal** or **Gspline**).

Mdistrib character specifying the prior distribution of the location of random effects. It is ignored if `hierar.center=FALSE`, in which case the location of the random effects is fixed to zero. It can be one of the following.

“fixed”

Locations of random effects are assumed to be fixed and are not updated.

“normal”

Locations of random effects are assumed to be a priori normally distributed. Parameters of the normal distribution are specified further by items `Mmean` and `Mvar`.

This is also a default choice when `Mdistrib` is not specified.

Mmean vector giving the prior means for the locations of random effects. It can be a single number only in which case it is recycled.

It is ignored when `hierar.center` is `FALSE` in which case all random effects have zero location.

Mvar vector giving the prior variances for the locations of random effects. It can be a single number only in which case it is recycled.

In the case that prior means and/or variances are different for different locations of random effects then they should be given in the similar order as specified above for argument `prior.fixed`.

It is ignored when `hierar.center` is `FALSE` in which case all random effects have zero location.

Ddistrib character specifying the prior distribution of the covariance matrix of random effects. It can be one of the following.

“fixed”

covariance matrix of random effects is assumed to be fixed and is not updated.

“wishart”

inverse of the covariance matrix of **normally** distributed random effects is assumed to have a priori Wishart distribution. Parameters of the Wishart distribution are specified further by items `Ddf` and `DinvScale`.

This is also a default choice when `Ddistrib` is not specified and random effects are normally distributed.

This option is not allowed when the random effects distribution is modelled using **G-splines**.

“sduniform”

when random effects are **normally** distributed then this option may be used when there is only a univariate random effect in the model. Its standard deviation is then assumed to follow a uniform distribution on the interval $(0, S)$. Value of S is specified further by an item `Dupper`.

When random effects distribution is modelled using **G-splines** then `sduniform`

can be used also for multivariate random effects. It is then assumed that the overall standard deviation τ_m in the m th margin follows a uniform distribution on the interval $(0, S_m)$.

“gamma”

when random effects are **normally** distributed then this option may be used when there is only a univariate random effect in the model. Its inverse variance is then assumed to have a Gamma prior with the shape specified further by the item `Dshape` and the rate (inverse scale) specified by the item `DinvScale`.

When random effects distribution is modelled using **G-splines** then gamma can be used also for multivariate random effects. It is then assumed that the overall inverse variance τ_m^{-2} in the m th margin follows a Gamma prior with the shapes specified further by the item `Dshape` and the rates (inverse scales) specified by the item `DinvScale`.

Ddf degrees of freedom ν_b for the Wishart prior distribution of the inverse covariance matrix of **normally** distributed random effects.

Dshape number or vector with shape parameters for the gamma priors of the variance components of the random effects.

If it is a single number and random effects are multivariate it may be recycled.

DinvScale number or matrix determining the inverse scale matrix S_b^{-1} for the Wishart prior distribution of the inverse covariance matrix of **normally** distributed random effects.

Or number or vector giving the rate (inverse scale) parameter(s) for the gamma priors of the variance components.

If it is a single number and `Ddistrib` is **wishart** then it is assumed that S_b^{-1} is diagonal with that single number on a diagonal.

Dupper upper limit for the uniform distribution of the standard deviation of the random effect(s) when `Ddistrib` is equal to **sduniform**.

If it is a single number and random effects are multivariate it may be recycled.

`prior.gspline` list specifying the G-spline distribution of random effects and prior distribution of the G-spline parameters. This argument is required only when `drandom` is equal to **gspline**. In the following let q denote the number (dimension) of random effects.

The list `prior.gspline` can have the following components.

K vector of length q or a number (it is recycled) which specifies, for each marginal G-spline, the number of knots on each side of the zero knot. That is, the m -th marginal G-spline has $2K_m + 1$ knots.

It is set to 15 if not explicitly specified.

delta vector of length q or a number (it is recycled) which specifies the distance between two consecutive knots for each marginal G-spline. That is, the m -th marginal G-spline has the following knots

$$\mu_{m,j} = j \delta_m, \quad j = -K_m, \dots, K_m.$$

It is set to 0.3 if not explicitly specified.

sigma vector of length q or a number (it is recycled) which specifies the basis standard deviation of each marginal G-spline.

σ_m is set to $(2/3)\delta_m$ if not explicitly specified.

CARorder vector of length q or a number (it is recycled) giving the order of the intrinsic conditional autoregression in the Gaussian Markov random field prior for the transformed G-spline weights in each margin.

It does not need to be specified when `neighbor.system` is different from `uniCAR`.

It is set to 3 if not explicitly specified.

neighbor.system character specifying the type of the Gaussian Markov random field in the prior for the transformed G-spline weights a of a **bivariate G-spline**. It does not have to be specified for univariate G-splines. It can be one of the following.

“uniCAR”

univariate (in each margin) conditional autoregression as described in Komárek and Lesaffre (2008). That is a priori

$$p(a | \lambda) \propto \exp \left[- \left\{ \frac{\lambda_1}{2} \sum_{j_1} \cdots \sum_{j_q} \left(\Delta_1^d a_{j_1, \dots, j_q} \right)^2 + \cdots + \frac{\lambda_q}{2} \sum_{j_1} \cdots \sum_{j_q} \left(\Delta_q^d a_{j_1, \dots, j_q} \right)^2 \right\} \right],$$

where Δ_m^d is a difference operator of order d in the m th margin, e.g.,

$$\Delta_1^3 a_{j_1, j_2, \dots, j_q} = a_{j_1, j_2, \dots, j_q} - 3a_{j_1-1, j_2, \dots, j_q} + 3a_{j_1-2, j_2, \dots, j_q} - a_{j_1-3, j_2, \dots, j_q},$$

and $\lambda = (\lambda_1, \dots, \lambda_q)'$ are smoothing hyperparameters.

This is also a default choice when `neighbor.system` is not specified.

“eight.neighbors”

this prior is applicable for **bivariate G-splines** only and is based on eight nearest neighbors in a spatial meaning. That is, except on edges, each full conditional of a depends only on eight nearest neighbors and local quadratic smoothing. The prior is then defined as

$$p(a | \lambda) \propto \exp \left\{ - \frac{\lambda}{2} \sum_{j_1=-K_1}^{K_1-1} \sum_{j_2=-K_2}^{K_2-1} (\Delta a_{j_1, j_2})^2 \right\},$$

where

$$\Delta a_{j_1, j_2} = a_{j_1, j_2} - a_{j_1+1, j_2} - a_{j_1, j_2+1} + a_{j_1+1, j_2+1}.$$

Parameter λ is a common smoothing hyperparameter.

“twelve.neighbors”

not (yet) implemented

Ldistrib character specifying the prior distribution of the smoothing hyperparameters λ_m , $m = 1, \dots, q$ (precision parameters of the Markov random fields in each margin) or of a common hyperparameter λ . It can be one of the following.

“fixed”

smoothing hyperparameters λ are fixed to their initial values and are not updated.

“gamma”

each of the smoothing hyperparameters $\lambda_1, \dots, \lambda_q$ is assumed to follow a Gamma prior with the shapes specified further by the item `Lshape` and the rates (inverse scales) specified further by the item `Linvscale`.

This is also a default choice when `Ldistrib` is not specified.

“sduniform”

square root of the inversion of each smoothing hyperparameter, i.e., $\sqrt{\lambda_1^{-1}}, \dots, \sqrt{\lambda_q^{-1}}$ is assumed to follow a priori a uniform distribution on the intervals $(0, S_m^\lambda)$.

Values of $S_1^\lambda, \dots, S_q^\lambda$ are specified further by the item `Lupper`.

Lequal logical indicating whether all smoothing hyperparameters should be kept equal.

It is set to `FALSE` if not explicitly specified and `neighbor.system` is `uniCAR`.

It is always `TRUE` when `neighbor.system` is different from `uniCAR`.

Lshape number or vector with shape parameters for the **gamma** priors of the smoothing hyperparameters λ .

If it is a single number and there is more than one smoothing hyperparameter λ in the model it may be recycled.

Linvscale number or vector with rate (inverse scale) parameters for the **gamma** priors of the smoothing hyperparameters λ .

If it is a single number and there is more than one smoothing hyperparameter λ in the model it may be recycled.

Lupper number or vector with upper limits for the uniform distribution on $\sqrt{\lambda^{-1}}$ parameters when `Ldistrib` is **sduniform**.

If it is a single number and there is more than one smoothing hyperparameter λ in the model it may be recycled.

Aident character specifying in which way the transformed G-spline weights (a coefficients) are identified. It can be one of the following.

“mean”

with this option, the a coefficients are forced to sum up to zero and have a zero mean.

Note that this option usually causes problems during MCMC, especially with bivariate G-splines. The reason is that if there are many almost zero weights, they lead to many negative a coefficients and to satisfy the zero mean constrain, there must be some a 's which are highly positive. When exponentiating them to get weights w , an overflow occur.

“reference”

with this option, one of the a coefficients in each margin is chosen as the reference one and is always equal to zero. Index of the reference coefficient is specified by the item `Areference` (see below).

This is also a default choice when `Aident` is not specified.

Areference vector or number (it is recycled) which specifies the index of the reference a coefficient in each margin in the case `Aident` is equal to **reference**. For the m -th margin, it must be an integer between $-K_m, \dots, K_m$. To avoid numerical problems, the index of the reference a coefficient may change during the MCMC.

AtypeUpdate character specifying in which way the transformed G-spline weights (a coefficients) are updated. It can be one of the following.

“slice”

slice sampler of Neal (2003).

“ars.quantile”

adaptive rejection sampling of Gilks and Wild (1992) with starting abscissae being quantiles of the envelop at the previous iteration.

“ars.mode”

adaptive rejection sampling of Gilks and Wild (1992) with starting abscissae being the mode plus/minus 3 times estimated standard deviation of the full conditional distribution.

“block”

all a coefficients are updated in 1 block using the Metropolis-Hastings algorithm. This is only available for univariate G-splines.

Default is `slice`.

`init.fixed` optional vector with initial values for fixed effects, supplied in the same order as described above in the argument `prior.fixed`.
If not given, initials are determined from the maximum-likelihood fit to the model where random effects are replaced by corresponding fixed effects.

`init.random` optional list with initial values for the random effects and parameters determining their distribution. It can have the following components.

b vector or matrix with initial values of cluster specific random effects. Number of rows of the matrix or the length of the vector must be equal to `length(unique(cluster))`. Columns of the matrix correspond to the random effects in the same order as described above for argument `prior.random`. If not given, initials are taken to be equal over clusters and equal to the corresponding fixed effects from the maximum-likelihood fit to the model with fixed effects only.

mean vector giving the initial values of the means of random effects.
If not given, initials are taken to be equal to the corresponding fixed effects from the maximum-likelihood fit to the model with fixed effects only.

var matrix giving the initial value for the covariance matrix of the random effects when `drandom` is **normal**.
Vector giving the initial values of marginal overall variances of the random effects when `drandom` is **gspline**.

`init.gspline` optional list with initial values related to the G-spline distribution of random effects (if `drandom` is equal to **gspline**). It can have the following components.

| | |
|-----------|--------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| | <p>lambda vector with initial values of the smoothing hyperparameters (precision parameters of the Markov random field). If not fixed and not given, the initials are sampled from the prior distribution.</p> <p>weights for univariate G-splines: vector with initial weights. For bivariate G-splines: matrix with initial weights. If the initial weights do not sum up to 1 they are re-scaled.</p> <p>alloc vector or matrix with initial component allocations for the individual random effects. For univariate random effects this should be a vector with numbers from $\{-K, \dots, K\}$. For bivariate random effects this should be a matrix with two columns where in the first column numbers from $\{-K_1, \dots, K_1\}$ appear and in the second column numbers from $\{-K_2, \dots, K_2\}$ appear.</p> |
| nsimul | <p>list indicating the length of the MCMC. It should have the following components.</p> <p>niter total number of the MCMC iterations after discarding the thinned values</p> <p>nthin thinning of the sample</p> <p>nburn length of the burn-in period</p> <p>nwrite frequency with which the iteration count changes. Further, during the burn-in, only every <code>nwrite</code>th sampled value is stored on the disk</p> |
| store | <p>list indicating which chains (out of these not stored by default) should be compulsory stored. The list has the logical components with the following names.</p> <p>prob if TRUE values of individual predictive probabilities are stored.</p> <p>b if TRUE values of cluster specific random effects are stored.</p> <p>alloc if TRUE values of allocation indicators are stored.</p> <p>acoef if TRUE and distribution of random effects is given as a bivariate G-spline values of log-G-spline weights (<i>a</i> coefficients) are stored for all components.</p> |
| dir | character string specifying the directory in which the sampled values are stored. |
| precision | precision with which the sampled values are written in files. |

Value

The function returns a complete list of parameters of the prior distribution and initial values.

The main task of this function is to sample from the posterior distribution using MCMC. Sampled values are stored in various files which are described below.

Files created

iteration.sim one column labeled `iteration` with indices of MCMC iterations to which the stored sampled values correspond.

betaF.sim sampled values of the fixed effects $\beta = (\beta^{(1)'}, \dots, \beta^{(C)'}, \beta^{*'})'$.

Note that in models with **G-spline** distributed random effects which are not hierarchically centered, the average effect of the covariates involved in the random effects (needed for inference) is obtained as a sum of the corresponding β coefficient and a scaled mean of the G-spline. β coefficients adjusted in this way are stored in the file `'betaRadj.sim'` (see below).

betaR.sim sampled values of the location parameters $\alpha = (\alpha^{(1)'}, \dots, \alpha^{(C)'}, \alpha^{*'})'$ of the random effects when the **hierarchical centering** was used.

Note that in models with **G-spline** distributed random effects which are hierarchically centered, the average effect of the covariates involved in the random effects (needed for inference) is obtained as a sum of the corresponding α coefficient and a mean of the G-spline. α coefficients adjusted in this way are stored in the file 'betaRadj.sim' (see below).

varR.sim variance components of the random effects. Format of the file depends on the assumed distribution of the random effects.

Normal random effects Let q be the dimension of the random effects. The first $0.5q(q+1)$ columns of 'varR.sim' contain a lower triangle (in column major order) of the matrix \mathbf{D}_b , the second $0.5q(q+1)$ columns of 'varR.sim' contain a lower triangle of the matrix \mathbf{D}_b^{-1} .

Univariate G-spline random effects The first column of 'varR.sim' contains the G-spline variance parameter τ^2 , the second column its inverse.

Bivariate G-spline random effects The first two columns of 'varR.sim' contain the G-spline variance parameters τ_1^2, τ_2^2 , the second two columns their inverse.

loglik.sim sampled values of the log-likelihood (conditioned by the values of random effects).

probability.sim sampled values of category probabilities for each observations.

Created only if store\$prob is TRUE.

b.sim sampled values of individual random effects.

Stores complete chains only if store\$b is TRUE.

Files created for models with G-spline distributed random effects

gspline.sim information concerning the fixed parameters of the G-spline which includes: dimension q of the G-spline, numbers of knots on each side of the reference knot for each margin (K_1, \dots, K_q) , basis standard deviations $\sigma_1, \dots, \sigma_q$ for each margin and knots $\mu_{1,-K_1}, \dots, \mu_{1,K_1}, \dots, \mu_{q,-K_q}, \dots, \mu_{q,K_q}$ for each margin.

weight.sim this file is created only for **bivariate** G-splines and stores the weights w of the G-spline which are higher than a certain threshold value. That is, the weights that are numerically equal to zero are not recorded here. The link between the weights and G-spline components is provided by the file 'knotInd.sim'.

knotInd.sim this file is created only for **bivariate** G-splines. In its first column, it stores the number of G-spline components for which the weights are recorded on a corresponding row of the file 'weight.sim'. Subsequently, it stores indices of the G-spline components for which the weights are given in the file 'weight.sim'. The indices are stored as single indices on the scale $0, \dots, (2K_1+1)(2K_2+1)-1$ such that index 0 corresponds to the component $(-K_1, -K_2)$, index 1 corresponds to the component $(-K_1+1, -K_2)$, ..., index K_1-1 corresponds to the component $(K_1, -K_2)$, etc.

logweight.sim sampled values of the transformed G-spline weights a .

For **univariate** G-spline, this file stores always complete chains, irrespective of the value of store\$acoef. For **bivariate** G-splines, this file stores complete chains only if store\$acoef is TRUE.

gmoment.sim first two moments of the unshifted and unscaled G-spline at each iteration.

For **univariate** G-spline the column labeled ‘gmean’ is equal to

$$\text{gmean} = \sum_{j=-K}^K w_j \mu_j$$

and the column labeled ‘gvar’ is equal to

$$\text{gvar} = \sum_{j=-K}^K w_j (\mu_j - \text{gmean})^2 + \sigma^2.$$

See Komárek and Lesaffre, E. (2008) for formulas that apply in the **bivariate** case.

Values stored here are the values of $\beta_1^*, \dots, \beta_q^*$ and $d_{1,1}^*, d_{2,1}^*, \dots, d_{q,q}^*$ as defined in Komárek and Lesaffre (2008).

betaRadj.sim sampled values of the average (overall) effects of the random effects. See notes under the files ‘betaF.sim’ and ‘betaR.sim’ above.

Values stored here are the values of $\gamma_1, \dots, \gamma_q$ as defined in Komárek and Lesaffre (2008).

varRadj.sim sampled components of the variance-covariance matrix of the random-effects.

Values stored here are the values of $d_{1,1}, d_{2,1}, \dots, d_{q,q}$ as defined in Komárek and Lesaffre (2008).

alloc.sim sampled values of the component allocations (in Komárek and Lesaffre (2008) denoted by r_i) for individual random effects.

For **univariate** G-spline, the allocations are stored on the scale $-K, \dots, K$.

For **bivariate** G-spline, the allocation are stored as single indeces on the scale $0, \dots, (2K_1 + 1)(2K_2 + 1) - 1$ where the link between the single and double indeces is the same as in the file ‘knotInd.sim’.

Stores complete chains only if store\$alloc is TRUE.

lambda.sim sampled values of smoothing hyperparameter(s) λ .

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

References

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See Also

[cumlogit](#), [logpoissonRE](#), [glm](#), [polr](#).

Examples

```
### See ex-Toenail.pdf and ex-Toenail.R
### available in the documentation
### to the package
```

cumlogitRE.predict *Prediction for logit and cumulative logit model with random effects*

Description

This function compute predictive probabilities of the response categories for specified combinations of covariates. It is based on the MCMC output obtained using [cumlogitRE](#).

Usage

```
cumlogitRE.predict(nobs, v, x, vb, xb, cluster,
  intcpt.random=FALSE, hierar.center=FALSE,
  drandom=c("normal", "gspline"),
  C=1, logit.order=c("decreasing", "increasing"),
  betaF, betaR, varR, is.varR=TRUE,
  prior.gspline,
  probs, values=FALSE,
  dir=getwd(), wfile, indfile, header=TRUE, logw, is.indfile,
  skip=0, nwrite)
```

Arguments

| | |
|------|------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| nobs | number of covariate combinations for which we want to perform a prediction |
| v | covariate combinations for which we want to perform a prediction. It should have the same structure as in cumlogitRE used to obtain the MCMC output |
| x | covariate combinations for which we want to perform a prediction. It should have the same structure as in cumlogitRE used to obtain the MCMC output |
| vb | covariate combinations for which we want to perform a prediction. It should have the same structure as in cumlogitRE used to obtain the MCMC output |
| xb | covariate combinations for which we want to perform a prediction. It should have the same structure as in cumlogitRE used to obtain the MCMC output |

| | |
|---------------|------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| cluster | vector defining pertinence of the single observations to clusters. It is useful when we want to predict longitudinal profiles. See also the same argument in cumlogitRE . |
| intcpt.random | see the same argument in cumlogitRE |
| hierar.center | see the same argument in cumlogitRE |
| drandom | see the same argument in cumlogitRE |
| C | see the same argument in cumlogitRE |
| logit.order | see the same argument in cumlogitRE |
| betaF | sampled values of the fixed effects. This should be a (sub)sample from the MCMC output stored in the file 'betaF.sim' |
| betaR | sampled values of the mean of random effects. This should be a (sub)sample from the MCMC output stored in the file 'betaR.sim' It is only needed if hierar.center is TRUE. |
| varR | sampled values of either (co)variance matrices or precision (matrices) for random effects if there are any. This should be a (sub)sample of either the first or second half of the columns stored in the file 'varR.sim' |
| is.varR | logical indicating whether varR gives (co)variance (is.varR TRUE) or precisions (inverse variances) (is.varR FALSE) |
| prior.gspline | if drandom is gspline this is a list specifying the G-splines. It should have the same structure as the same argument in cumlogitRE used to obtain the MCMC output. However, it is satisfactory if the items K , delta and sigma are given. |
| probs | probabilities for which the (pointwise) sample quantiles of the predictive probabilities should be computed. If not given only average (and values) of the predictive probabilities are computed |
| values | if TRUE also values of the predictive probabilities at each (MCMC) iteration are returned. If FALSE only sample mean (and quantiles) of the predictive probabilities are returned |
| dir | character giving the directory where the file with (sampled) G-spline (log-)weights is stored. Needed only if drandom is gspline . |
| wfile | character giving the name of the file with (sampled) G-spline (log-)weights. Needed only if drandom is gspline . In most cases, for univariate G-spline this argument will be equal to "logweight.sim" and for bivariate G-spline equal to "weight.sim". |
| indfile | character giving the name of the file where we stored indices of these G-spline components for which the weights are stored in the file given by wfile. The corresponding file should have the same structure as 'knotInd.sim' created by cumlogitRE . Needed only if is.indfile is TRUE. In most cases, for univariate G-spline it does not have to be specified and for bivariate G-spline it will be equal to "knotInd.sim". |

| | |
|------------|----------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| header | logical indicating whether the files wfile, indfile contain a header. Needed only if drandom is gspline . |
| logw | logical indicating whether the file wfile contains logarithms of the weights. Needed only if drandom is gspline . In most cases, for univariate G-spline it will be TRUE and for bivariate G-spline it will be FALSE. |
| is.indfile | logical. If TRUE then wfile contains only the non-zero weights and the G-spline is reconstructed using indfile. If FALSE then wfile must contain on each row weights of all components and indfile is ignored. Needed only if drandom is gspline and random effects are bivariate. |
| skip | number of data rows that should be skipped at the beginning of the files wfile, indfile. |
| nwrite | frequency with which is the user informed about the progress of computation (every nwriteth iteration count of iterations change) |

Value

A list with the following components (description below applies for the case with prob=0.5)

| | |
|--------|-----------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| Mean | a matrix with $C + 1$ columns giving in each row posterior predictive mean of category probabilities $P(Y = 0), \dots, P(Y = C)$ for a given covariate combination. |
| 50% | a matrix with $C + 1$ columns giving in each row posterior predictive quantile (here 50% quantile) of category probabilities for a given covariate combination. There is one component of this type in the resulting list for each value of probs. |
| values | a matrix with $(C + 1)n$ columns, where n denotes the number of covariate combinations for which we perform the prediction, and number of rows equal to the length of the MCMC. The first $C + 1$ columns give sampled category probabilities for the first covariate combination, the second $C + 1$ columns give sampled category probabilities for the second covariate combination etc. It is returned only if values is TRUE. |

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

References

Komárek, A. and Lesaffre, E. (2008). Generalized linear mixed model with a penalized Gaussian mixture as a random-effects distribution. *Computational Statistics and Data Analysis*, **52**, 3441–3458.

See Also

[cumlogitRE](#), [cumlogit](#), [glm](#), [polr](#).

`densplotAK`*Probability density function estimate from MCMC output*

Description

Displays a plot of the density estimate for each variable in `x`, calculated by the density function.

This is slightly modified version of `densplot` function of a `coda` package to conform to my personal preferences.

Usage

```
densplotAK(x, plot=TRUE, show.obs=FALSE, bwf, bty="n", main="",  
           xlim, ylim, xlab, ylab, ...)
```

Arguments

| | |
|-------------------------------------|--------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| <code>x</code> | an <code>mcmc</code> or <code>mcmc.list</code> object. |
| <code>plot</code> | if TRUE this function works more or less in the same way as <code>coda</code> function <code>densplot</code> function. If FALSE this function returns one data frame for each chain with computed density which can be used for future plotting. |
| <code>show.obs</code> | show observations along the x-axis? |
| <code>bwf</code> | function for calculating the bandwidth. If omitted, the bandwidth is calculate by 1.06 times the minimum of the standard deviation and the interquartile range divided by 1.34 times the sample size to the negative one fifth power. |
| <code>xlim, ylim, xlab, ylab</code> | further arguments passed to the <code>plot.default</code> function. |
| <code>bty, main, ...</code> | further arguments passed to the <code>plot.default</code> function. |

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

See Also

[densplot](#).

epileptic

Epileptic seizures

Description

This data set considers information from a clinical trial of 59 epileptics, reported by Thall and Vail (1990). For each patient, the number of epileptic seizures was recorded during a baseline period of eight weeks. Patients were then randomized to treatment with the anti-epileptic drug progabide, or to placebo in addition to standard chemotherapy. The number of seizures was then recorded in four consecutive two-weeks intervals.

Usage

```
data(epileptic)
```

Format

A data frame with 295 observations on the following 5 variables.

`id` a unique identifier for the subject in the study.

`seizure` a numeric vector giving the number of epileptic seizures.

`visit` a numeric vector giving the number of the visit, 0=baseline, and 1,2,3, and 4 for the four consecutive two-weeks intervals.

`trt` a numeric vector giving the treatment group.

`age` a numeric vector giving the age at the entry.

Source

Thall, P. F., and Vail, S. C. (1990) Some covariance models for longitudinal count data with overdispersion. *Biometrics*, **46**, 657–671.

References

Breslow, N. E. and Clayton, D. G. (1993). Approximate inference in generalized linear mixed models. *Journal of the American Statistical Association*, **88**, 9–25.

Kleinman, K. P. and Ibrahim, J. G. (1998). A semi-parametric Bayesian approach to generalized linear mixed models. *Statistics in Medicine*, **17**, 2579–2596.

See Also

[epilepticBC](#)

Examples

```
data(epileptic)
## maybe str(epileptic); plot(epileptic) ...
```

 epilepticBC

Epileptic seizures, Breslow and Clayton transformed

Description

This data set considers information from a clinical trial of 59 epileptics, reported by Thall and Vail (1990). For each patient, the number of epileptic seizures was recorded during a baseline period of eight weeks. Patients were then randomized to treatment with the anti-epileptic drug progabide, or to placebo in addition to standard chemotherapy. The number of seizures was then recorded in four consecutive two-weeks intervals.

This is a transformed version of the original data (`epileptic`) which allows directly to fit models described in Breslow and Clayton (1993) or in Kleinman and Ibrahim (1998).

There are 4 rows in the dataset for each of 59 patients.

Usage

```
data(epilepticBC)
```

Format

A data frame with 236 observations on the following 10 variables.

`id` a unique identifier for the subject in the study.

`visit` a numeric vector giving the number of the visit, 1,2,3, and 4 for the four two-weeks intervals following the baseline measurement.

`seizure0` a numeric vector giving the number of epileptic seizures in the eight-week period prior the randomization.

`age` a numeric vector giving the age (in years) at the entry.

`Seizure` a numeric vector giving the number of epileptic seizures in a given two-week interval.

`Base` a numeric vector giving the transformed `seizure0`. $\text{Base} = \log(\text{seizure0}/4)$.

`Trt` a numeric vector giving the treatment group.

`Base.Trt` a numeric vector giving the interaction covariate $\text{Base} * \text{Trt}$.

`Age` a numeric vector giving the transformed age. $\text{Age} = \log(\text{age})$.

`Visit` a numeric vector giving the centered visit. $\text{Visit} = (2 * \text{visit} - 5) / 10$.

Source

Thall, P. F., and Vail, S. C. (1990) Some covariance models for longitudinal count data with overdispersion. *Biometrics*, **46**, 657–671.

References

Breslow, N. E. and Clayton, D. G. (1993). Approximate inference in generalized linear mixed models. *Journal of the American Statistical Association*, **88**, 9–25.

Kleinman, K. P. and Ibrahim, J. G. (1998). A semi-parametric Bayesian approach to generalized linear mixed models. *Statistics in Medicine*, **17**, 2579–2596.

See Also[epileptic](#)**Examples**

```
data(epilepticBC)
## maybe str(epilepticBC); plot(epilepticBC) ...
```

| | |
|-------------------|------------------------------------------------------------|
| glmmAK.files2coda | <i>Conversion of sampled values into coda mcmc objects</i> |
|-------------------|------------------------------------------------------------|

Description

It takes the values sampled by [cumlogitRE](#) or [logpoissonRE](#) and stored in *.sim files and converts them into coda [mcmc](#) objects.

Usage

```
glmmAK.files2coda(dir, drandom=c("none", "normal", "gspline"),
  quiet=FALSE, skip=0,
  params=list(prob=FALSE, ecount=FALSE, b=FALSE, alloc=FALSE))
```

Arguments

| | |
|---------|-------------------------------------------------------------------------------------------------------------------------------------------|
| dir | character specifying a directory with sampled values |
| drandom | string specifying the distribution of random effects used in the original cumlogitRE or logpoissonRE call |
| quiet | logical, passed to scan function |
| skip | number of MCMC iterations that should be skipped at the beginning of the chain |
| params | a list of logical values specifying which of the optional sampled parameters should be read |

Value

A list of coda [mcmc](#) objects.

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

See Also

[mcmc](#).

Description

Moments, density and random generation for the Gaussian Markov random field with mean equal to 'mean', precision matrix equal to 'Q' (or covariance matrix equal to 'Sigma') and possibly constrained by a linear constraint 'Ax=b'.

Generation of random numbers is performed by Algorithm 2.6 in Rue and Held (2005, pp. 38).

Usage

```
momentsGMRF(mean=0, Q=1, Sigma, A, b=0)
rGMRF(n, mean=0, Q=1, Sigma, A, b=0)
dGMRF(x, mean=0, Q=1, Sigma, A, b=0, log=FALSE)
dGMRF2(x, mean=0, Q=1, Sigma, A, b=0, log=FALSE)
```

Arguments

| | |
|-------|--------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| mean | vector of mean. If length(mean) is equal to 1, it is recycled and all components have the same mean. |
| Q | precision matrix of the GMRF. |
| Sigma | covariance matrix of the GMRF. Only one of Q and Sigma must be given. If Sigma is supplied, precision is computed from Σ as $Q = \Sigma^{-1}$. |
| A | optional matrix defining the constraint $Ax = b$ for sampled vectors x . If not supplied, the GMRF is assumed to be unconstrained. Currently at most 1 constraint is allowed, that is A must be a vector and b a number. |
| b | vector or the right-hand side of the constraint. If length(b) is equal to 1, it is recycled and all constraint right-hand sides are the same. |
| n | number of observations to be sampled. |
| x | vector or matrix of the points where the density should be evaluated. |
| log | logical; if TRUE, log-density is computed |

Value

Some objects.

Value for momentsGMRF

A list with the components:

mean mean of the (constrained) GMRF

Sigma covariance matrix of the (constrained) GMRF

and the following attributes:

mean.unconstr mean of the GMRF before imposing the constraints

Sigma.unconstr covariance matrix Σ_u of the GMRF before imposing the constraints

Q.unconstr.cholesky Cholesky decomposition of the matrix $Q_u = \Sigma_u^{-1}$

nconstraint number of constraints

A left-hand side of the constraints

b right-hand side of the constraints

Value for rGMRF

A list with the components:

x vector or matrix with sampled values

log.dens vector with the values of the log-density evaluated in the sampled values

Value for dGMRF, dGMRF2

A vector with evaluated values of the (log-)density

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

References

Rue, H. and Held, L. (2005). *Gaussian Markov Random Fields: Theory and Applications*. Boca Raton: Chapman and Hall/CRC.

See Also

[dnorm](#), [Mvnorm](#).

Examples

```
set.seed(1977)

mu <- c(0, 6, 8)
L <- matrix(1:9, nrow=3)
L[upper.tri(L, diag=FALSE)] <- 0
Sigma <- L %*% t(L)
Q <- chol2inv(chol(Sigma))
```

```

A <- rep(1, nrow(Sigma))
b <- 0

##### Unconstrained GMRF
##### =====
## Moments
momentsGMRF(mean=mu, Sigma=Sigma)
momentsGMRF(mean=mu, Q=Q)

## Random numbers
z <- rGMRF(1000, mean=mu, Sigma=Sigma)
apply(z$x, 2, mean)
var(z$x)

## Random numbers, again
z <- rGMRF(10, mean=mu, Sigma=Sigma)
print(z)

## Values of the log-density
dGMRF(z$x, mean=mu, Sigma=Sigma, log=TRUE)
dGMRF(z$x, mean=mu, Q=Q, log=TRUE)
dGMRF2(z$x, mean=mu, Sigma=Sigma, log=TRUE)
dGMRF2(z$x, mean=mu, Q=Q, log=TRUE)

## Values of the density
dGMRF(z$x, mean=mu, Sigma=Sigma)
dGMRF(z$x, mean=mu, Q=Q)
dGMRF2(z$x, mean=mu, Sigma=Sigma)
dGMRF2(z$x, mean=mu, Q=Q)

##### Constrained GMRF
##### =====
## Moments
momentsGMRF(mean=mu, Sigma=Sigma, A=A, b=b)
momentsGMRF(mean=mu, Q=Q, A=A, b=b)

## Random numbers
z <- rGMRF(1000, mean=mu, Sigma=Sigma, A=A, b=b)
apply(z$x, 2, mean)
var(z$x)

## Random numbers, again
z <- rGMRF(10, mean=mu, Sigma=Sigma, A=A, b=b)
print(z)
A %% t(z$x)

## Values of the log-density
dGMRF(z$x, mean=mu, Sigma=Sigma, A=A, b=b, log=TRUE)
dGMRF(z$x, mean=mu, Q=Q, A=A, b=b, log=TRUE)
dGMRF2(z$x, mean=mu, Sigma=Sigma, A=A, b=b, log=TRUE)
dGMRF2(z$x, mean=mu, Q=Q, A=A, b=b, log=TRUE)

```

```
## Values of the log-density
dGMRF(z$x, mean=mu, Sigma=Sigma, A=A, b=b)
dGMRF(z$x, mean=mu, Q=Q, A=A, b=b)
dGMRF2(z$x, mean=mu, Sigma=Sigma, A=A, b=b)
dGMRF2(z$x, mean=mu, Q=Q, A=A, b=b)
```

| | |
|----------|-----------------------------------------------------------------------------------------------------|
| gspline1 | <i>Density and random number generation from a univariate G-spline (penalized Gaussian mixture)</i> |
|----------|-----------------------------------------------------------------------------------------------------|

Description

Univariate G-spline (penalized Gaussian mixture) is distributed as

$$\alpha + \sum_{j=-K}^K w_j \mathbf{N}(\tau \mu_j, \tau^2 \sigma_j^2)$$

Usage

```
rgspline1(n, mu, sigma, weight, intcpt=0, scale=1, logw=TRUE)
```

```
dgspline1(x, mu, sigma, weight, intcpt=0, scale=1, logw=TRUE)
```

Arguments

| | |
|--------|---------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| n | number of observations to be generated |
| x | grid of values at which we evaluate the G-spline values |
| mu | a vector with G-spline knots μ_j ($j = -K, \dots, K$) (means of basis G-splines) |
| sigma | basis standard deviation(s) σ_j ($j = -K, \dots, K$). If a single number is supplied then it is assumed that all basis G-splines have the same standard deviation. Alternatively a vector of the same length as mu can be given in which case the basis G-splines do not necessarily have the same standard deviations |
| weight | a vector with G-spline (log-)weights. It should have the same length as mu |
| intcpt | G-spline intercept value α |
| scale | G-spline scale value τ |
| logw | logical indicating whether logarithmic weights are supplied in weight |

Value

Values of the density or generated random numbers.

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

Examples

```

knots <- c(-2, 0, 2)
sigma <- 1
weight <- c(0.3, 0.1, 0.6)
intcpt <- 3
scale <- 0.2
xgrid <- seq(1.8, 4.2, length=300)

dx <- dgspline1(xgrid, mu=knots, sigma=sigma, weight=weight,
  intcpt=intcpt, scale=scale, logw=FALSE)
x <- rgspline1(100, mu=knots, sigma=sigma, weight=weight,
  intcpt=intcpt, scale=scale, logw=FALSE)
hist(x, col="seagreen2", prob=TRUE, xlim=range(xgrid), xlab="x", ylab="g(x)")
lines(xgrid, dx, col="red", lwd=2)

```

gspline2

Density and random number generation from a bivariate G-spline (penalized Gaussian mixture)

Description

Density has not been implemented yet.

Bivariate G-spline (penalized Gaussian mixture) is distributed as

$$(\alpha_1, \alpha_2)' + \sum_{j_1=-K_1}^{K_1} \sum_{j_2=-K_2}^{K_2} w_{j_1, j_2} \mathbf{N}_2((\tau_1 \mu_{1, j_1}, \tau_2 \mu_{2, j_2})', \text{diag}(\tau_1^2 \sigma_{1, j_1}^2, \tau_2^2 \sigma_{2, j_2}^2))$$

Usage

```

rgspline2(n, mu1, mu2, sigma1, sigma2, weight, knotInd,
  intcpt=0, scale=1, logw=TRUE)

```

Arguments

| | |
|--------|--------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| n | number of observations to be generated |
| mu1 | a vector with G-spline knots μ_{1, j_1} ($j_1 = -K_1, \dots, K_1$) (means of basis G-splines) in the 1st margin |
| mu2 | a vector with G-spline knots μ_{2, j_2} ($j_2 = -K_2, \dots, K_2$) (means of basis G-splines) in the 2nd margin |
| sigma1 | basis standard deviation(s) σ_{1, j_1} ($j_1 = -K_1, \dots, K_1$) in the 1st margin. If a single number is supplied then it is assumed that all basis G-splines have the same standard deviation. Alternatively a vector of the same length as mu1 can be given in which case the basis G-splines do not necessarily have the same standard deviations |

| | |
|---------|----------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| sigma2 | basis standard deviation(s) σ_{2,j_2} ($j_2 = -K_2, \dots, K_2$) in the 2nd margin. If a single number is supplied then it is assumed that all basis G-splines have the same standard deviation. Alternatively a vector of the same length as mu2 can be given in which case the basis G-splines do not necessarily have the same standard deviations |
| weight | a vector or matrix with G-spline (log-)weights. If <code>missing(knotInd)</code> then it should have the same length as <code>length(mu1)*length(mu2)</code> . If <code>knotInd</code> is supplied then <code>weight</code> must be a vector of the same length as <code>knotInd</code> which contains (log-)weights of G-spline components having non-zero weights. |
| knotInd | If supplied then it contains indices of G-spline components which correspond to non-zero weights. Indices should be on the scale from 0 to <code>length(mu1)*length(mu2)-1</code> (similarly like in the file <code>knotInd.sim</code> created, e.g., by the function <code>cumlogitRE</code>) |
| intcpt | G-spline intercept value(s) $(\alpha_1, \alpha_2)'$ If a single value is supplied then it is assumed that intercept values in both margins are the same. |
| scale | G-spline scale value(s) $(\tau_1, \tau_2)'$ If a single value is supplied then it is assumed that scale values in both margins are the same. |
| logw | logical indicating whether logarithmic weights are supplied in <code>weight</code> |

Value

Values of the density or generated random numbers.

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

Examples

```
knots1 <- c(-2, 0, 2)
knots2 <- c(-2, -1, 0, 1, 2)
sigma1 <- 0.5
sigma2 <- 0.5
intcpt <- c(3, -1)
scale <- c(0.2, 0.5)

weight <- matrix(c(1,2,1, 2,3,2, 3,4,3, 2,3,2, 1,2,1), ncol=5)
xA <- rgspline2(1000, mu1=knots1, mu2=knots2, sigma1=sigma1,
  sigma2=sigma2, weight=weight, intcpt=intcpt, scale=scale, logw=FALSE)
xB <- rgspline2(1000, mu1=knots1, mu2=knots2, sigma1=sigma1,
  sigma2=sigma2, weight=log(weight), intcpt=intcpt, scale=scale, logw=TRUE)

oldpar <- par(mfrow=c(1, 2), bty="n")
plot(xA[,2], xA[,1], pch=16, col="red")
abline(h=intcpt[1]+scale[1]*knots1, col="orange")
abline(v=intcpt[2]+scale[2]*knots2, col="orange")
```

```

plot(xB[,2], xB[,1], pch=16, col="blue")
abline(h=intcpt[1]+scale[1]*knots1, col="orange")
abline(v=intcpt[2]+scale[2]*knots2, col="orange")
par(oldpar)

### Only selected components with non-zero weights
logweight2 <- c(-1, -0.5, 1)
knotInd <- c(0, 5, 14)
yA <- rgspline2(1000, mu1=knots1, mu2=knots2, sigma1=sigma1,
  sigma2=sigma2, weight=logweight2, knotInd=knotInd,
  intcpt=intcpt, scale=scale, logw=TRUE)
yB <- rgspline2(1000, mu1=knots1, mu2=knots2, sigma1=sigma1,
  sigma2=sigma2, weight=exp(logweight2), knotInd=knotInd,
  intcpt=intcpt, scale=scale, logw=FALSE)
oldpar <- par(mfrow=c(1, 2), bty="n")
plot(yA[,2], yA[,1], pch=16, col="red")
abline(h=intcpt[1]+scale[1]*knots1, col="orange")
abline(v=intcpt[2]+scale[2]*knots2, col="orange")
plot(yB[,2], yB[,1], pch=16, col="blue")
abline(h=intcpt[1]+scale[1]*knots1, col="orange")
abline(v=intcpt[2]+scale[2]*knots2, col="orange")
par(oldpar)

logweight3 <- log(weight)[-c(5,8,11)]
knotInd3 <- c(0,1,2, 3,5, 6,8, 9,11, 12,13,14)
zA <- rgspline2(1000, mu1=knots1, mu2=knots2, sigma1=sigma1,
  sigma2=sigma2, weight=logweight3, knotInd=knotInd3,
  intcpt=intcpt, scale=scale, logw=TRUE)
zB <- rgspline2(1000, mu1=knots1, mu2=knots2, sigma1=sigma1,
  sigma2=sigma2, weight=exp(logweight3), knotInd=knotInd3,
  intcpt=intcpt, scale=scale, logw=FALSE)
oldpar <- par(mfrow=c(1, 2), bty="n")
plot(zA[,2], zA[,1], pch=16, col="red")
abline(h=intcpt[1]+scale[1]*knots1, col="orange")
abline(v=intcpt[2]+scale[2]*knots2, col="orange")
plot(zB[,2], zB[,1], pch=16, col="blue")
abline(h=intcpt[1]+scale[1]*knots1, col="orange")
abline(v=intcpt[2]+scale[2]*knots2, col="orange")
par(oldpar)

```

Description

Fits the poisson log-linear regression model using the maximum-likelihood. The log-likelihood is maximized using the Newton-Raphson algorithm (the same as Fisher scoring in this case). The function returns the inverse of the observed and expected information matrix.

Usage

```
logpoisson(y, x, offset=0, epsilon=1e-08, maxit=25, trace=FALSE)

## S3 method for class 'logpoisson'
print(x, ...)

## S3 method for class 'logpoisson'
summary(object, ...)
```

Arguments

y response vector taking integer values or zero.
x matrix or data.frame with covarites.
 Intercept is included by default in the model and should not be included in x.
offset possible offset term. It is assumed to be equal to zero if not specified.
epsilon positive convergence tolerance ε . The iterations converge when

$$\left| \frac{\ell_{new} - \ell_{old}}{\ell_{new}} \right| \leq \varepsilon,$$

where ℓ denotes the value of the log-likelihood.

maxit integer giving the maximal number of iterations.
trace logical indicating if output should be produced for each iteration.
object an object of class "logpoisson".
... other arguments passed to print or summary.

Value

An object of class "logpoisson". This has components

coefficients the coefficients of the linear predictor.
loglik the value of the log-likelihood.
score the score vector.
vcov the inverse of the information matrix.
linear.predictors the values of the linear predictor for each observation.
fitted.values the values of fitted counts for each observation.
converged logical indicating whether the optimization routine converged.
iter number of iterations performed
y
x

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

References

Agresti, A. (2002). *Categorical Data Analysis. Second edition.* Hoboken: John Wiley & Sons. Section 7.2.

See Also

[glm](#).

Examples

```
set.seed(1977)
n <- 100
x1 <- rbinom(n, 1, 0.4)
x2 <- runif(n, 0, 1)
eta <- 5 + 0.1*x1 - 0.2*x2
mu <- exp(eta)
y <- rpois(n, mu)

### Fit the model using poisson
Xmat <- data.frame(x1=x1, x2=x2)
fit <- logpoisson(y=y, x=Xmat)
summary(fit)

### Fit the model using standard glm
fit0 <- glm(y~x1+x2, family=poisson(link="log"))
summary(fit0)
```

logpoissonRE

Poisson log-linear regression with random effects

Description

This function implements MCMC sampling for the Poisson log-linear model. Details are given in Komárek and Lesaffre (2008). On as many places as possible, the same notation as in this paper is used also in this manual page.

In general, the following log-linear model for response Y is assumed:

$$\log\{E(Y)\} = \eta,$$

where the form of the linear predictor η depends on whether a hierarchical centering is used or not. In the following, β denotes fixed effects and b random effects.

No hierarchical centering (DEFAULT)

The linear predictor has the following form

$$\eta = \beta'(x', x'_b) + b'x_b,$$

where b is a vector of random effects with zero location.

Hierarchical centering

The linear predictor has the following form

$$\eta = \beta'x + b'x_b,$$

b is a vector of random effects with location α .

For description of the rest of the model, see [cumlogitRE](#).

Usage

```
logpoissonRE(y, x, xb, offset=0, cluster,
  intcpt.random=FALSE,
  hierar.center=FALSE,
  drandom=c("normal", "gspline"),
  prior.fixed,
  prior.random,
  prior.gspline,
  init.fixed,
  init.random,
  init.gspline,
  nsimul = list(niter=10, nthin=1, nburn=0, nwrite=10),
  store = list(ecount=FALSE, b=FALSE, alloc=FALSE, acoef=FALSE),
  dir=getwd(),
  precision=8)
```

Arguments

| | |
|----------------------------|----------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| <code>y</code> | response vector taking integer values or zero. |
| <code>x</code> | vector, matrix or data.frame with covariates for fixed effects. |
| <code>xb</code> | vector, matrix or data.frame with covariates for random effects. If you want to include random intercept , do it by setting the argument <code>intcpt.random</code> to TRUE. The intercept column should not be included in <code>xb</code> . |
| <code>offset</code> | optional vector of the offset term. |
| <code>cluster</code> | see cumlogitRE . |
| <code>intcpt.random</code> | see cumlogitRE . |
| <code>hierar.center</code> | see cumlogitRE . |
| <code>drandom</code> | see cumlogitRE . |
| <code>prior.fixed</code> | see cumlogitRE . |
| <code>prior.random</code> | see cumlogitRE . |
| <code>prior.gspline</code> | see cumlogitRE . |
| <code>init.fixed</code> | see cumlogitRE . |
| <code>init.random</code> | see cumlogitRE . |
| <code>init.gspline</code> | see cumlogitRE . |
| <code>nsimul</code> | see cumlogitRE . |

| | |
|-----------|----------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| store | list indicating which chains (out of these not stored by default) should be compulsory stored. The list has the logical components with the following names. ecount if TRUE values of individual predictive (expected) counts are stored. b if TRUE values of cluster specific random effects are stored. alloc if TRUE values of allocation indicators are stored. acoef if TRUE and distribution of random effects is given as a bivariate G-spline values of log-G-spline weights (a coefficients) are stored for all components. |
| dir | see cumlogitRE . |
| precision | see cumlogitRE . |

Value

See [cumlogitRE](#).

Files created

iteration.sim see [cumlogitRE](#).

betaF.sim sampled values of the fixed effects β .

Note that in models with **G-spline** distributed random effects which are not hierarchically centered, the average effect of the covariates involved in the random effects (needed for inference) is obtained as a sum of the corresponding β coefficient and a scaled mean of the G-spline. β coefficients adjusted in this way are stored in the file 'betaRadj.sim' (see below).

betaR.sim sampled values of the location parameters α of the random effects when the **hierarchical centering** was used.

Note that in models with **G-spline** distributed random effects which are hierarchically centered, the average effect of the covariates involved in the random effects (needed for inference) is obtained as a sum of the corresponding α coefficient and a mean of the G-spline. α coefficients adjusted in this way are stored in the file 'betaRadj.sim' (see below).

varR.sim see [cumlogitRE](#).

loglik.sim see [cumlogitRE](#).

expectcount.sim sampled values of predictive (expected) counts for each observations.

Created only if store\$ecount is TRUE.

b.sim see [cumlogitRE](#).

Files created for models with G-spline distributed random effects

See [cumlogitRE](#).

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

References

- Agresti, A. (2002). *Categorical Data Analysis. Second edition*. Hoboken: John Wiley & Sons.
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- Neal, R. M. (2003). Slice sampling (with Discussion). *The Annals of Statistics*, **31**, 705–767.
- Komárek, A. and Lesaffre, E. (2008). Generalized linear mixed model with a penalized Gaussian mixture as a random-effects distribution. *Computational Statistics and Data Analysis*, **52**, 3441–3458.
- Molenberghs, G. and Verbeke, G. (2005). *Models for Discrete Longitudinal Data*. New York: Springer Science+Business Media.

See Also

[logpoisson](#), [cumlogitRE](#), [poisson](#), [glm](#).

Examples

```
### See ex-Epileptic.pdf and ex-Epileptic.R
### available in the documentation
### to the package
```

logpoissonRE.predict *Prediction for Poisson log-linear regression with random effects*

Description

This function compute predictive (expected) counts for specified combinations of covariates. It is based on the MCMC output obtained using [logpoissonRE](#).

Usage

```
logpoissonRE.predict(nobs, x, xb, offset, cluster,
  intcpt.random=FALSE, hierar.center=FALSE,
  drandom=c("normal", "gspline"),
  betaF, betaR, varR, is.varR=TRUE,
  prior.gspline,
  probs, values=FALSE,
  dir=getwd(), wfile, indfile, header=TRUE, log, is.indfile,
  skip=0, nwrite)
```

Arguments

| | |
|---------------|-------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| nobs | number of covariate combinations for which we want to perform a prediction |
| x | covariate combinations for which we want to perform a prediction. It should have the same structure as in logpoissonRE used to obtain the MCMC output |
| xb | covariate combinations for which we want to perform a prediction. It should have the same structure as in logpoissonRE used to obtain the MCMC output |
| offset | optional offset vector. |
| cluster | vector defining pertinence of the single observations to clusters. It is useful when we want to predict longitudinal profiles. See also the same argument in logpoissonRE . |
| intcpt.random | see the same argument in logpoissonRE |
| hierar.center | see the same argument in logpoissonRE |
| drandom | see the same argument in logpoissonRE |
| betaF | sampled values of the fixed effects. This should be a (sub)sample from the MCMC output stored in the file 'betaF.sim' |
| betaR | sampled values of the mean of random effects. This should be a (sub)sample from the MCMC output stored in the file 'betaR.sim' It is only needed if hierar.center is TRUE. |
| varR | sampled values of either (co)variance matrices or precision (matrices) for random effects if there are any. This should be a (sub)sample of either the first or second half of the columns stored in the file 'varR.sim' |
| is.varR | logical indicating whether varR gives (co)variance (is.varR TRUE) or precisions (inverse variances) (is.varR FALSE) |
| prior.gspline | if drandom is gspline this is a list specifying the G-splines. It should have the same structure as the same argument in logpoissonRE used to obtain the MCMC output. However, it is satisfactory if the items K , delta and sigma are given. |
| probs | probabilities for which the (pointwise) sample quantiles of the predictive counts should be computed. If not given only average (and values) of the predictive counts are computed |
| values | if TRUE also values of the predictive counts at each (MCMC) iteration are returned. If FALSE only sample mean (and quantiles) of the predictive probabilities are returned |
| dir | character giving the directory where the file with (sampled) G-spline (log-)weights is stored. Needed only if drandom is gspline . |
| wfile | character giving the name of the file with (sampled) G-spline (log-)weights. Needed only if drandom is gspline . In most cases, for univariate G-spline this argument will be equal to "logweight.sim" and for bivariate G-spline equal to "weight.sim". |

| | |
|-----------|-------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| infile | character giving the name of the file where we stored indices of these G-spline components for which the weights are stored in the file given by wfile. The corresponding file should have the same structure as 'knotInd.sim' created by logpoissonRE . Needed only if is.infile is TRUE. In most cases, for univariate G-spline it does not have to be specified and for bivariate G-spline it will be equal to "knotInd.sim". |
| header | logical indicating whether the files wfile, infile contain a header. Needed only if drandom is gspline . |
| logw | logical indicating whether the file wfile contains logarithms of the weights. Needed only if drandom is gspline . In most cases, for univariate G-spline it will be TRUE and for bivariate G-spline it will be FALSE. |
| is.infile | logical. If TRUE then wfile contains only the non-zero weights and the G-spline is reconstructed using infile. If FALSE then wfile must contain on each row weights of all components and infile is ignored. Needed only if drandom is gspline and random effects are bivariate. |
| skip | number of data rows that should be skipped at the beginning of the files wfile, infile. |
| nwrite | frequency with which is the user informed about the progress of computation (every nwrite iteration count of iterations change) |

Value

A list with the following components (description below applies for the case with prob=0.5)

| | |
|--------|--------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| Mean | a matrix with 1 column giving in each row posterior predictive mean of the count $E(Y)$ for a given covariate combination. |
| 50% | a matrix with 1 column giving in each row posterior predictive quantile (here 50% quantile) of the count for a given covariate combination. There is one component of this type in the resulting list for each value of probs. |
| values | a matrix with n columns, where n denotes the number of covariate combinations for which we perform the prediction, and number of rows equal to the length of the MCMC. Each column gives sampled counts a given covariate combination. It is returned only if values is TRUE. |

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

References

Komárek, A. and Lesaffre, E. (2008). Generalized linear mixed model with a penalized Gaussian mixture as a random-effects distribution. *Computational Statistics and Data Analysis*, **52**, 3441–3458.

See Also

[logpoissonRE](#), [logpoisson](#), [glm](#).

maxPosterProb

G-spline utility

Description

For given G-spline basis and given data, it determines for each data point the G-spline component for which the value of the basis density is maximal.

Usage

```
maxPosterProb(data, intercept, std.dev, K, delta, sigma)
```

Arguments

- | | |
|-----------|------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| data | numeric vector or matrix with data. If given as a matrix then rows correspond to observations and columns to margins. |
| intercept | numeric vector of length 1 or <code>ncol(data)</code> with intercepts for each margin. If given as a number, it is recycled. |
| std.dev | numeric vector of length 1 or <code>ncol(data)</code> with standard deviations for each margin. If given as a number, it is recycled. |
| K | numeric vector of length 1 or <code>ncol(data)</code> which specifies, for each marginal G-spline, then number of knots on each side of the zero knot. That is, the i -th marginal G-spline has $2K_i + 1$ knots. If given as a number, it is recycled. |
| delta | numeric vector of length 1 or <code>ncol(data)</code> which specifies the distance between two consecutive knots for each marginal G-spline. That is, the i -th marginal G-spline has the following knots $\mu_{i,j} = j \delta_i, j = -K_i, \dots, K_i.$ If given as a number, it is recycled. |
| sigma | numeric vector of length 1 or <code>ncol(data)</code> with basis standard deviations for marginal G-splines. If given as a number, it is recycled. |

Value

Matrix which specifies determined components (indexces are on scale $-K_i, \dots, K_i$).

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

See Also

cumlogitRE, logpoissonRE.

Examples

```
N <- 100
intcpt <- c(0, 5, 15)
std.dev <- c(1, 0.5, 3)
data <- data.frame(b1=rnorm(N, intcpt[1], std.dev[1]),
                  b2=rnorm(N, intcpt[2], std.dev[2]),
                  b3=rnorm(N, intcpt[3], std.dev[3]))
alloc <- maxPosterProb(data=data, intercept=intcpt, std.dev=std.dev,
                      K=15, delta=0.3, sigma=0.2)

par(mfrow=c(1, 3), bty="n")
for (i in 1:3) hist(alloc[,i], prob=TRUE, col="seagreen3",
                  xlab="Allocation", breaks=(-15):15,
                  main=paste("Margin ", i, sep=""))
```

QuantileFun

Sample quantiles

Description

This is (almost) the same as `quantile(x)`, or `apply(x, 1, quantile)` or `apply(x, 2, quantile)`.
The motivation to write it was to validate my C++ function.

Usage

```
QuantileFun(x, probs=seq(0, 1, 0.25), vals.in.cols=TRUE)
```

Arguments

| | |
|---------------------------|------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| <code>x</code> | values of the function |
| <code>probs</code> | numeric vector of probabilities with values in $[0, 1]$ |
| <code>vals.in.cols</code> | if TRUE then it is assumed that function f evaluated in a specific grid point over (MCMC) iterations is stored in a column of <code>x</code> . That is (MCMC) iterations correspond to rows. |

Value

A data.frame with 1 row for each probs value.

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

See Also

[quantile](#).

Examples

```
probs <- c(0, 0.25, 0.354, 0.5, 0.75, 1)

x <- rnorm(1001)
QuantileFun(x, probs=probs)
quantile(x, probs=probs)

n <- 1001
xx <- data.frame(x1=rnorm(n), x2=rgamma(n, shape=1, rate=1), x3=1:n)
QuantileFun(xx, probs=probs, vals.in.cols=TRUE)
apply(xx, 2, quantile, probs=probs)

xx2 <- t(xx)
QuantileFun(xx2, probs=probs, vals.in.cols=FALSE)
apply(xx2, 1, quantile, probs=probs)
```

 scanFH

Read Data Values

Description

Read numeric data into a data frame from a file. Header is assumed to be present in the file.

Usage

```
scanFH(file, quiet=FALSE)
```

Arguments

file the name of a file to read data values from. If the specified file is "", then input is taken from the keyboard (or stdin if input is redirected). (In this case input can be terminated by a blank line or an EOF signal, Ctrl-D on Unix and Ctrl-Z on Windows.)

Otherwise, the file name is interpreted *relative* to the current working directory (given by `getwd()`), unless it specifies an *absolute* path. Tilde-expansion is performed where supported.

Alternatively, `file` can be a [connection](#), which will be opened if necessary, and if so closed at the end of the function call. Whatever mode the connection is opened in, any of LF, CRLF or CR will be accepted as the EOL marker for a line and so will match `sep = "\n"`.

`file` can also be a complete URL.

To read a data file not in the current encoding (for example a Latin-1 file in a UTF-8 locale or conversely) use a [file](#) connection setting the encoding argument.

quiet logical: if FALSE (default), scan() will print a line, saying how many items have been read.

Details

See [scan](#).

Value

data.frame with read data values.

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

References

Becker, R. A., Chambers, J. M. and Wilks, A. R. (1988) *The New S Language*. Wadsworth & Brooks/Cole.

See Also

[scan](#)

Examples

```
cat("x y z", "1 2 3", "1 4 6", "10 20 30", file="ex.data", sep="\n")
pp <- scanFH("ex.data", quiet=FALSE)
pp <- scanFH("ex.data", quiet= TRUE)
print(pp)
unlink("ex.data") # tidy up
```

summaryGspline1

Summary for a univariate G-spline (penalized Gaussian mixture)

Description

This function is primarily designed to work out the MCMC output from functions [cumlogitRE](#) and [logpoissonRE](#) in which a distribution of the univariate random effect was specified as a G-spline. It computes posterior pointwise mean and quantiles for a G-spline density based on the MCMC output.

Usage

```
summaryGspline1(x, mu, sigma,
  standard=TRUE, intcpt, scale,
  probs, values=FALSE,
  dir=getwd(), wfile="logweight.sim", header=TRUE, logw=TRUE,
  skip=0, nwrite)
```

Arguments

| | |
|----------|-----------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| x | grid of values at which we want to evaluate the G-spline density |
| mu | a vector with G-spline knots (means of basis G-splines) |
| sigma | basis standard deviation(s). If a single number is supplied then it is assumed that all basis G-splines have the same standard deviation. Alternatively a vector of the same length as mu can be given in which case the basis G-splines do not necessarily have the same standard deviations |
| standard | logical, if TRUE then the standardized (zero-mean, unit-variance) G-spline densities are computed and summarized |
| intcpt | a vector with sampled intercept values. If not supplied it is assumed that all intercepts are equal to zero. It does not have to be supplied if standard=TRUE. |
| scale | a vector with sampled values of the G-spline scale (in most of my papers denoted by τ) If not supplied it is assumed that all scale values are equal to one. It does not have to be supplied if standard=TRUE. |
| probs | probabilities for which the (pointwise) sample quantiles of the G-spline density should be computed. If not given only average (and values) of the G-spline are computed |
| values | if TRUE also values of the G-spline at each (MCMC) iteration are returned. If FALSE only sample mean (and quantiles) of the G-spline are returned |
| dir | character giving the directory where the file with (sampled) G-spline (log-)weights is stored |
| wfile | character giving the name of the file with (sampled) G-spline (log-)weights |
| header | logical indicating whether the file wfile contains a header |
| logw | logical indicating whether the file wfile contains logarithms of the weights |
| skip | number of data rows that should be skipped at the beginning of the file wfile |
| nwrite | frequency with which is the user informed about the progress of computation (every nwrite iteration count of iterations change) |

Value

A list with the following components (component 'values' is present only when the argument values was TRUE):

| | |
|---------|--------------------------------------------------------------------------------------------------------------------|
| summary | a data.frame with the following columns (the description below applies to the situation when probs=0.5): |
| 'x' | a grid of values at which the G-spline density is evaluated |
| 'Mean' | pointwise posterior mean of the G-spline density |
| '50%' | pointwise posterior 50% quantile of the G-spline density. There is one column of this type for each probs value |
| values | a matrix with one column for each x value and number of rows equal to the |

length of the MCMC. In each row, there is a G-spline density evaluated at one MCMC iteration.

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

Examples

```
### See ex-Toenail.pdf, ex-Toenail.R
### and ex-Epileptic.pdf, ex-Epileptic.R
### available in the documentation
### to the package
```

summaryGspline2

Summary for a bivariate G-spline (penalized Gaussian mixture)

Description

This function is primarily designed to work out the MCMC output from functions `cumlogitRE` and `logpoissonRE` in which a distribution of the bivariate random effect was specified as a G-spline. It computes posterior pointwise mean and quantiles for a G-spline density based on the MCMC output.

Besides the summary for the joint bivariate G-spline density it also directly computes summaries for both marginal G-splines.

Usage

```
summaryGspline2(x1, x2, mu1, mu2, sigma1, sigma2,
  standard=TRUE, intcpt, scale,
  probs, values=FALSE,
  dir=getwd(), wfile="weight.sim", indfile="knotInd.sim",
  header=TRUE, logw=FALSE, is.indfile=TRUE,
  skip=0, nwrite)
```

Arguments

| | |
|--------|-------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| x1 | grid of values for the first margin at which we evaluate the G-spline density |
| x2 | grid of values for the second margin at which we evaluate the G-spline density |
| mu1 | a vector with G-spline knots (means of basis G-splines) for the first margin |
| mu2 | a vector with G-spline knots (means of basis G-splines) for the second margin |
| sigma1 | basis standard deviation(s) for the first margin. If a single number is supplied then it is assumed that all basis G-splines in the first margin have the same standard deviation. Alternatively a vector of the same length as mu1 can be given in which case the basis G-splines in the first margin do not necessarily have the same standard deviations |

| | |
|------------|----------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| sigma2 | basis standard deviation(s) for the second margin. If a single number is supplied then it is assumed that all basis G-splines in the second margin have the same standard deviation. Alternatively a vector of the same length as mu2 can be given in which case the basis G-splines in the second margin do not necessarily have the same standard deviations |
| standard | if TRUE then the standardized (zero-mean, unit-variance) G-splines are computed and summarized |
| intcpt | a two-column matrix with sampled intercept values. If not supplied it is assumed that all intercepts are equal to zero. It does not have to be supplied if standard=TRUE. |
| scale | a two-column matrix with sampled intercept values of the G-spline scale (in most of my papers denoted by τ) If not supplied it is assumed that all scale values are equal to one. It does not have to be supplied if standard=TRUE. |
| probs | probabilities for which the (pointwise) sample quantiles of the G-spline should be computed. If not given only average (and values) of the G-spline are computed |
| values | if TRUE also values of the G-spline at each (MCMC) iteration are returned. If FALSE only sample mean (and quantiles) of the G-spline are returned |
| dir | character giving the directory where the file with (sampled) G-spline (log-)weights is stored |
| wfile | character giving the name of the file with (sampled) G-spline (log-)weights |
| indfile | character giving the name of the file where it is indicated which G-spline components correspond to non-zero weights. It does not have to be supplied if is.indfile=FALSE. |
| header | logical indicating whether the files wfile, indfile contain a header |
| logw | logical indicating whether the file wfile contains logarithms of the weights |
| is.indfile | logical. If TRUE then wfile contains only the non-zero weights and the G-spline is reconstructed using indfile. If FALSE then wfile must contain on each row weights of all components and indfile is ignored. |
| skip | number of data rows that should be skipped at the beginning of the files wfile, indfile |
| nwrite | frequency with which is the user informed about the progress of computation (every nwrite iteration count of iterations change) |

Value

A list with the following components (component 'values', 'values1', 'values2' are present only when the argument values was TRUE). The description below applies to the situation when probs=0.5.

summary a list with the components

| | |
|----------|----------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| 'x1' | grid of values for the first margin at which we evaluate the G-spline density |
| 'x2' | grid of values for the second margin at which we evaluate the G-spline density |
| 'Mean' | a matrix with <code>length(x1)</code> rows and <code>length(x2)</code> columns giving the pointwise posterior mean of the joint G-spline density |
| '50%' | a matrix with <code>length(x1)</code> rows and <code>length(x2)</code> columns giving the pointwise posterior 50% quantile of the joint G-spline density. There is a matrix of this type for each probs value. |
| summary1 | a <code>data.frame</code> with the following columns |
| 'x' | a grid of values at which the first marginal G-spline density is evaluated |
| 'Mean' | pointwise posterior mean of the first marginal G-spline density |
| '50%' | pointwise posterior 50% quantile of the first marginal G-spline density. There is one column of this type for each probs value |
| summary2 | a <code>data.frame</code> for the second marginal G-spline having the same structure as 'summary1'. |
| values | a matrix with one column for each (x1, x2) value and number of rows equal to the length of the MCMC. In each row, there is a joint G-spline density evaluated at one MCMC iteration. |
| values1 | a matrix with one column for each x1 value and number of rows equal to the length of the MCMC. In each row, there is the first marginal G-spline density evaluated at one MCMC iteration. |
| values1 | a matrix with one column for each x1 value and number of rows equal to the length of the MCMC. In each row, there is the second marginal G-spline density evaluated at one MCMC iteration. |

Author(s)

Arnošt Komárek <arnost.komarek[AT]mff.cuni.cz>

Examples

```
### See ex-Toenail.pdf, ex-Toenail.R
### and ex-Epileptic.pdf, ex-Epileptic.R
### available in the documentation
### to the package
```

Description

This data set considers information from a longitudinal clinical trial in dermatology which was set up to compare the efficacy of two oral treatments for toenail infection (De Backer et al., 1998). One of the end points of the study was the degree of onycholysis which expresses the degree of separation of the nail plate from the nail-bed (0, absent; 1, mild; 2, moderate; 3, severe) and was evaluated at seven visits (approximately on weeks 0, 4, 8, 12, 24, 36 and 48). In total, 1 908 measurements on 294 patients are available. In this dataset, only a dichotomized onycholysis (0, absent or mild; 1, moderate or severe) is given (variable infect).

IMPORTANT NOTICE: The data have kindly been made available by Novartis, Belgium. The source of the data must be acknowledged in any publication which uses them (see Lesaffre and Spiessens, 2001 for more details).

Usage

```
data(toenail)
```

Format

A data frame with 1908 observations on the following 5 variables.

```
idnr  identification number of the patient
infect binary response
trt   treatment group
time  time of measurement (in months)
visit visit number
```

Source

Lesaffre, E. and Spiessens, B. (2001). On the effect of the number of quadrature points in a logistic random-effects model: An example. *Applied Statistics*, **50**, 325–335.

References

De Backer, M., De Vroey, C., Lesaffre, E., Scheys, I., and De Keyser, P. (1998). Twelve weeks of continuous onychomycosis caused by dermatophytes: A double blind comparative trial of terbafine 250 mg/day versus itraconazole 200 mg/day. *Journal of the American Academy of Dermatology*, **38**, S57–S63.

Examples

```
data(toenail)
## maybe str(toenail); plot(toenail) ...
```

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